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Application of Probabilistic Method on Daehee Sequences

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Abstract. In this paper, we investigate some combinatorial sequences based on Daehee and Changhee numbers and polynomials, then derive their moment representations in use of probabilistic method. We also provide identities related to Daehee numbers, derangement numbers, Cauchy numbers of the second kind, and Stirling numbers of the first kind.

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1. Introduction and Preliminaries

Throughout this paper, we use the following notations:

$$\mathbb{N} = \{1, 2, 3, \dots\}, \mathbb{Z}_{\geq 0} = \{0, 1, 2, \dots\}.$$

Let $D_{n,\xi}^{(k)}(x)$ denote the *n*th twisted Daehee polynomials of order $k \in \mathbb{N}$, which are defined by the generating function [2] to be

$$\left(\frac{\ln(1+\xi t)}{\xi t}\right)^k (1+\xi t)^x = \sum_{n=0}^{\infty} D_{n,\xi}^{(k)}(x) \frac{t^n}{n!}.$$
 (1)

In special case, when x=0, $D_{n,\xi}^{(k)}=D_{n,\xi}^{(k)}(0)$ are called twisted Daehee numbers of order k. Similarly, $D_n^{(k)}=D_{n,1}^{(k)}$ are higher-order Daehee numbers, $D_{n,\xi}=D_{n,\xi}^{(1)}$ are twisted Daehee numbers, and $D_n=D_{n,1}^{(1)}$ are Daehee numbers.

Let $\hat{D}_{n,\xi}^{(k)}(x)$ denote the *n*th twisted Daehee polynomials of the second kind of order $k \in \mathbb{N}$, which are defined by the generating function[2] to be

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$$\left(\frac{(1+\xi t)ln(1+\xi t)}{\xi t}\right)^k (1+\xi t)^x = \sum_{n=0}^{\infty} \hat{D}_{n,\xi}^{(k)}(x) \frac{t^n}{n!}.$$
 (2)

In special case, when x=0, $\hat{D}_{n,\xi}^{(k)}=\hat{D}_{n,\xi}^{(k)}(0)$ are called higher-order twisted Daehee numbers of the second kind. Similarly, $\hat{D}_n^{(k)}=\hat{D}_{n,1}^{(k)}$ are higher-order Daehee numbers of the second kind, $\hat{D}_{n,\xi}=\hat{D}_{n,\xi}^{(1)}$ are twisted Daehee numbers of the second kind, and $\hat{D}_n=\hat{D}_{n,1}^{(1)}$ are Daehee numbers of the second kind.

Let $Ch_n^{(r)}(x)$ denote the nth Changhee polynomials of order $r \in \mathbb{N}$, which are defined by the generating function[1] to be

$$\left(\frac{2}{2+t}\right)^r (1+t)^x = \sum_{n=0}^{\infty} Ch_n^{(r)}(x) \frac{t^n}{n!}.$$
 (3)

When x = 0, $Ch_n^{(r)}(0) = Ch_n^{(r)}$ are called higher-order Changhee numbers.

Let $\hat{C}h_n^{(r)}(x)$ denote the *n*th Changhee polynomials of order $r \in \mathbb{N}$ of the second kind, which are defined by the generating function[1] to be

$$\left(\frac{2}{2+t}\right)^r (1+t)^{x+r} = \sum_{n=0}^{\infty} \hat{C}h_n^{(r)}(x) \frac{t^n}{n!}.$$
 (4)

When x = 0, $\hat{C}h_n^{(r)}(0) = \hat{C}h_n^{(r)}$ are called higher-order Changhee numbers of the second kind.

Remark 1. [see 7] If f and g are exponential generating functions, and

$$fg = \left(\sum_{r=0}^{\infty} \frac{a_r x^r}{r!}\right) \left(\sum_{s=0}^{\infty} \frac{b_s x^s}{s!}\right),$$

then the coefficients of $\frac{x^n}{n!}$ in fg are given by

$$\left[\frac{x^n}{n!}\right](fg) = \sum_{r=0}^n \binom{n}{r} a_r b_{n-r}.$$

Remark 2. Throughout this paper, symbol E denotes the expectation operator defined by

$$Ef(X) = \int_{-\infty}^{+\infty} f(x)p(x)dx,$$

where random variable X is continuous, whose density function is p(x). Specially, when $f(x) = x^n$, EX^n denotes n-order moment of random variable X.

(i) When
$$r.v \ u \sim U[0,1], Eu^n = \frac{1}{n+1},$$

(ii) When $r.v X \sim \Gamma(1,1)$, $EX^n = n!$.

Definition 1. The characteristic function of random variable X is defined as

$$\varphi(t) = Ee^{itX}, i^2 = -1, -\infty < t < \infty.$$
(5)

When the moments of all orders of r.v. X exist, the following relation expression holds true,

$$EX^{n} = \left[\frac{(it)^{n}}{n!}\right]\varphi(t), \quad i^{2} = -1. \tag{6}$$

Remark 3. [see 5] If random variable X is distributed as $\Gamma(\alpha, \lambda)$, where $\alpha, \lambda > 0$, its characteristic function is

$$\varphi(t) = Ee^{itX} = (1 - \frac{it}{\lambda})^{-\alpha}. (7)$$

Remark 4. X and Y are two random variables, when Cov(X,Y)=0, we have $E(XY)=EX \cdot EY$, where

$$Cov(X,Y) = E(XY) - E(X)E(Y)$$

Then we give three lemmas to introduce moment representations of some special combinatorial sequences.

Lemma 1. [see 6] Assume that $r.v X \sim \Gamma(u, 1)$, with $u \sim U[0, 1]$ is a random variable that follows uniform distribution, and X, u are independent respectively, then Cauchy numbers of the second kind \hat{C}_n , whose generating function is

$$\sum_{n=0}^{\infty} \hat{C}_n \frac{x^n}{n!} = \frac{x}{(1+x)\ln(1+x)},\tag{8}$$

have the following moment representation,

$$\hat{C}_n = EX^n, \quad n \geqslant 0. \tag{9}$$

Lemma 2. [see 6] Assume that $r.v \ X \sim \Gamma(1,1)$, then derangement numbers $d_n = n! \sum_{k=0}^{n} \frac{(-1)^k}{k!}$ have the following moment representation,

$$d_n = E(X-1)^n, \quad n \geqslant 0. \tag{10}$$

Lemma 3. [see 6] Suppose that $r.v.s.u_1, u_2, \cdots, i.i.d \sim U[0,1], r.v.s. \Gamma_1, \Gamma_2, \cdots, i.i.d \sim \Gamma(1,1), r.v.u_i$ and Γ_j are independent respectively for all i,j. When $n,k \geq 1$, Stirling numbers of the first kind s(n,k) satisfy

$$s(n,k) = (-1)^{n-k} \binom{n}{k} E(u_1 \Gamma_1 + u_2 \Gamma_2 + \dots + u_k \Gamma_k)^{n-k}.$$
 (11)

It is demanded that s(n,0)=s(0,k)=0, s(0,0)=1.

2. Moment Representations of Daehee and Changhee Sequences

In this section, we use probabilistic method to derive moment representations about two kinds of higher-order twisted Daehee numbers and polynomials, and two kinds of higher-order Changhee numbers and polynomials.

Theorem 1. Assume that r.v.s $u_1, u_2, ..., i.i.d \sim U[0,1], \Gamma_1, \Gamma_2, ..., i.i.d \sim \Gamma(1,1),$ and for all $i, j, r.v u_i$ and Γ_j are independent. When $n, m - k \in \mathbb{Z}_{\geq 0}$, $k \in \mathbb{N}$, we have

$$D_{n,\xi}^{(k)} = (-\xi)^n E(u_1 \Gamma_1 + u_2 \Gamma_2 + \dots + u_k \Gamma_k)^n,$$
(12)

$$D_{m-k,\xi}^{(k)} = (-\xi)^{m-k} \frac{k}{m} E(u_1 \Gamma_1 + \dots + u_{k-1} \Gamma_{k-1} + \Gamma_k)^{m-k}.$$
 (13)

Proof. The generating function of higher-order twisted Daehee numbers is known as

$$\sum_{n=0}^{\infty} D_{n,\xi}^{(k)} \frac{t^n}{n!} = \left(\frac{\ln(1+\xi t)}{\xi t}\right)^k,\tag{14}$$

Taking the coefficients of t^n in the left-hand side of Eq.(14), we get

$$\frac{D_{n,\xi}^{(k)}}{n!} = [t^n] (\frac{\ln(1+\xi t)}{\xi t})^k = [(-t)^n] (\sum_{i\geq 0} \frac{(\xi t)^i}{i+1})^k = [(-t)^n] (\sum_{i\geq 0} t^i E(\xi u)^i)^k, \tag{15}$$

$$(\sum_{i\geqslant 0} t^{i} E(\xi u)^{i})^{k} = \sum_{n=0}^{\infty} \sum_{i_{1}+\dots+i_{k}=n} (E(\xi u_{1})^{i_{1}}) \cdots (E(\xi u_{k})^{i_{k}}) t^{n}$$

$$= \sum_{n=0}^{\infty} \frac{1}{n!} \sum_{i_{1}+\dots+i_{k}=n} \binom{n}{i_{1}, i_{2}, \dots i_{k}} (E(\xi u_{1})^{i_{1}}) \cdots (E(\xi u_{k})^{i_{k}}) (i_{1}!) \cdots (i_{k}!) t^{n}$$

$$= \sum_{n=0}^{\infty} \frac{1}{n!} \sum_{i_{1}+\dots+i_{k}=n} \binom{n}{i_{1}, i_{2}, \dots i_{k}} (E(\xi u_{1})^{i_{1}}) \cdots (E(\xi u_{k})^{i_{k}}) (E\Gamma_{1}^{i_{1}}) \cdots (E\Gamma_{k}^{i_{k}}) t^{n}$$

$$= \sum_{n=0}^{\infty} \frac{\xi^{n}}{n!} E[\sum_{i_{1}+\dots+i_{k}=n} \binom{n}{i_{1}, i_{2}, \dots i_{k}} (u_{1}\Gamma_{1})^{i_{1}} \cdots (u_{k}\Gamma_{k})^{i_{k}}] t^{n}$$

$$= \sum_{n=0}^{\infty} \frac{\xi^{n}}{n!} E(u_{1}\Gamma_{1}+\dots+u_{k}\Gamma_{k})^{n} t^{n}. \tag{16}$$

From Eq.(15) and Eq.(16), we can see that

$$\frac{D_{n,\xi}^{(k)}}{n!} = \frac{(-\xi)^n}{n!} E(u_1 \Gamma_1 + \dots + u_k \Gamma_k)^n,$$

thus we have

$$D_{n,\xi}^{(k)} = (-\xi)^n E(u_1 \Gamma_1 + \dots + u_k \Gamma_k)^n.$$

Eq.(13) can be proved by the following equation[see 6]

$$E(u_1\Gamma_1 + u_2\Gamma_2 + \dots + u_k\Gamma_k)^{m-k} = \frac{k}{m}E(u_1\Gamma_1 + \dots + u_{k-1}\Gamma_{k-1} + \Gamma_k)^{m-k}$$
 (17)

Corollary 1. In theorem 1, when $\xi = 1$, we obtain the moment representation of higher-order Daehee numbers[3]

$$D_n^{(k)} = (-1)^n E(u_1 \Gamma_1 + \dots + u_k \Gamma_k)^n.$$
(18)

Corollary 2. In theorem 1, when k = 1, we obtain the moment representation of twisted Daehee numbers[2]

$$D_{n,\xi} = (-\xi)^n E(u\Gamma)^n. \tag{19}$$

Corollary 3. In theorem1, when $\xi = 1, k = 1$, we obtain the moment representation of Daehee numbers[4]

$$D_n = (-1)^n E(u\Gamma)^n. (20)$$

Theorem 2. Suppose that $r.v\ u \sim U[0,1],\ \Gamma \sim \Gamma(1,1),\ r.v\ u$ and Γ are independent, then twisted Daehee numbers of the second kind of order k satisfy

$$\hat{D}_{n,\xi}^{(k)} = \sum_{i=0}^{n} \binom{n}{i} \xi^{n-i}(k)_{n-i} D_{i,\xi}^{(k)}.$$
 (21)

Proof. The generating function of $\hat{D}_{n,\xi}^{(k)}$ is given by

$$\left(\frac{\ln(1+\xi t)}{\xi t}(1+\xi t)\right)^k = \sum_{n=0}^{\infty} \hat{D}_{n,\xi}^{(k)} \frac{t^n}{n!}.$$
 (22)

From theorem 1, the left-hand side of Eq.(22) can be written as

$$\left(\frac{\ln(1+\xi t)}{\xi t}\right)^{k} (1+\xi t)^{k}
= \sum_{n=0}^{\infty} (-\xi)^{n} E(u_{1}\Gamma_{1} + u_{2}\Gamma_{2} + \dots + u_{k}\Gamma_{k})^{n} \frac{t^{n}}{n!} \sum_{n=0}^{\infty} (k)_{n} \frac{(\xi t)^{n}}{n!}
= \sum_{n=0}^{\infty} \sum_{i=0}^{n} \binom{n}{i} (-\xi)^{i} E(u_{1}\Gamma_{1} + u_{2}\Gamma_{2} + \dots + u_{k}\Gamma_{k})^{i} (k)_{n-i} \xi^{n-i} \frac{t^{n}}{n!}
= \sum_{n=0}^{\infty} \sum_{i=0}^{n} \binom{n}{i} \xi^{n-i} (k)_{n-i} D_{i,\xi}^{(k)} \frac{t^{n}}{n!},$$

where $(k)_n = k(k-1)\cdots(k-n+1)$. By comparing the coefficients of $\frac{t^n}{n!}$, theorem 2 is proved.

Corollary 4. In theorem 2, taking $k = 1, n \ge 1$, we obtain the moment form of twisted Daehee numbers of the second kind,

$$\hat{D}_{n,\xi} = \frac{(-1)^{n-1}}{n} \xi^n E(u\Gamma)^n, \ n \geqslant 1, \ \hat{D}_{0,\xi} = 1.$$
 (23)

Corollary 5. From corollary 2 and corollary 4, when $n \ge 1$, the following relationship holds true,

$$-n\hat{D}_{n,\xi} = D_{n,\xi}, \quad n \geqslant 1. \tag{24}$$

Theorem 3. Under the circumstance of theorem 1, higher-order twisted Daehee polynomials $D_{n,\mathcal{E}}^{(k)}(x)$ have the following moment representation3

$$D_{n,\xi}^{(k)}(x) = \sum_{i=0}^{n} \binom{n}{i} \xi^{n-i}(x)_{n-i} D_{i,\xi}^{(k)}.$$
 (25)

Proof. Proof of theorem 3 is similar to the one of theorem 2.

Theorem 4. Suppose that r.v.s $u_1, u_2, ..., i.i.d \sim U[0,1], \Gamma_1, \Gamma_2, ..., i.i.d \sim \Gamma(1,1),$ $X \sim \Gamma[-x, \frac{1}{\xi}], (x < 0, \xi > 0), and r.v u_i, \Gamma_j and X are independent for all i, j, then higher-order twisted Daehee polynomials <math>D_{n,\xi}^{(k)}(x)$ satisfy

$$D_{n,\xi}^{(k)}(x) = E[\xi(u_1\Gamma_1 + \dots + u_k\Gamma_k) + X]^n.$$
 (26)

Proof. Replacing t by -it in the generating function of higher-order twisted Daehee polynomials, and according to remark 3 and 4, we have

$$\sum_{n=0}^{\infty} D_{n,\xi}^{(k)}(x) \frac{(-it)^n}{n!} = \left(\frac{\ln(1-\xi it)}{-\xi it}\right)^k (1-\xi it)^{-(-x)}$$

$$= \sum_{n=0}^{\infty} (-1)^n \xi^n E(u_1 \Gamma_1 + \dots + u_k \Gamma_k)^n \frac{(-it)^n}{n!} \sum_{n=0}^{\infty} EX^n \frac{(it)^n}{n!}$$

$$= \sum_{n=0}^{\infty} \sum_{i=0}^n \binom{n}{i} \xi^i E(u_1 \Gamma_1 + \dots + u_k \Gamma_k)^i EX^{n-i} \frac{(it)^n}{n!}$$

$$= \sum_{n=0}^{\infty} E[\xi(u_1 \Gamma_1 + \dots + u_k \Gamma_k) + X]^n \frac{it^n}{n!}.$$

By comparing the coefficients of $\frac{(it)^n}{n!}$, we obtain theorem 4.

Theorem 5. Suppose that r.v $u \sim U[0,1]$, $\Gamma \sim \Gamma(1,1)$, r.v u and Γ are independent respectively, twisted Daehee polynomials of the second kind of order k satisfy

$$\hat{D}_{n,\xi}^{(k)}(x) = \sum_{i=0}^{n} \binom{n}{i} \xi^{n} (-1)^{i} E(u_{1} \Gamma_{1} + \dots + u_{k} \Gamma_{k})^{i} (x+k)_{n-i}.$$
 (27)

Proof. Proof of theorem 5 is similar to that of theorem 2.

Theorem 6. Suppose that r.v.s $u_1, u_2, ..., i.i.d \sim U[0,1], \Gamma_1, \Gamma_2, ..., i.i.d \sim \Gamma(1,1),$ $X \sim \Gamma[-x-k,\frac{1}{\xi}], (x<-k,\xi>0),$ and r.v u_i, Γ_j and X are independent for all i, j, then higher-order twisted Daehee polynomials of the second kind $\hat{D}_{n,\xi}^{(k)}(x)$ satisfy

$$\hat{D}_{n,\xi}^{(k)}(x) = (-1)^n E[\xi(u_1\Gamma_1 + \dots + u_k\Gamma_k) + X]^n.$$
(28)

Proof. Proof of theorem 6 is similar to the one of theorem 4.

Theorem 7. Assume r.v $X \sim \Gamma[r, 2]$, then higher-order Changhee numbers satisfy

$$Ch_n^{(r)} = (-1)^n EX^n.$$
 (29)

Proof. Replacing t by -it, where $i^2 = -1$ in the generating function of $Ch_n^{(r)}$, we have

$$\left(\frac{2}{2-it}\right)^r = \sum_{n=0}^{\infty} Ch_n^{(r)} \frac{(-it)^n}{n!},\tag{30}$$

The left-hand side of Eq.(30) can be written as

$$(1 - \frac{it}{2})^{-r} = Ee^{itX} = \sum_{n=0}^{\infty} EX^n \frac{(it)^n}{n!},$$
(31)

from Eq.(30) and Eq.(31), we obtain

$$\sum_{n=0}^{\infty} Ch_n^{(r)} \frac{(-it)^n}{n!} = \sum_{n=0}^{\infty} (-1)^n EX^n \frac{(-it)^n}{n!}.$$

By comparing the coefficients of $\frac{(-it)^n}{n!}$, theorem 7 is proved.

Theorem 8. Under the circumstance of theorem 7, higher-order Changhee polynomials of the second kind $\hat{C}h_n^{(r)}(x)$ satisfy

$$\hat{C}h_n^{(r)}(x) = \sum_{k=0}^n \binom{n}{k} (-1)^k (x+r)_{n-k} EX^k.$$
(32)

Proof. The generating function of $\hat{C}h_n^{(r)}(x)$ can be written as

$$\begin{split} \sum_{n=0}^{\infty} \hat{C} h_n^{(r)}(x) \frac{t^n}{n!} &= (\frac{2}{2+t})^r (1+t)^{x+r} \\ &= \sum_{n=0}^{\infty} (-1)^n E X^n \frac{t^n}{n!} \sum_{n=0}^{\infty} (x+r)_n \frac{t^n}{n!} \end{split}$$

$$= \sum_{n=0}^{\infty} \sum_{k=0}^{n} \binom{n}{k} (-1)^k EX^k (x+r)_{n-k} \frac{t^n}{n!}.$$

By comparing the coefficients of $\frac{t^n}{n!}$, theorem 8 is proved.

Corollary 6. In theorem 8, taking x = 0, we obtain the moment form of higher-order Changhee numbers of the second kind,

$$\hat{C}h_n^{(r)} = \sum_{k=0}^n \binom{n}{k} (-1)^k (r)_{n-k} EX^k.$$
(33)

Corollary 7. Since the generating function of higher-order Changhee polynomials is

$$\sum_{n=0}^{\infty} Ch_n^{(r)}(x) \frac{t^n}{n!} = (\frac{2}{t+2})^r (1+t)^x.$$
 (34)

From theorem 8, we can see that the moment form of $Ch_n^{(r)}(x)$ is

$$Ch_n^{(r)}(x) = \sum_{k=0}^n \binom{n}{k} (-1)^k (x)_{n-k} EX^k.$$
 (35)

3. Identities of Daehee Numbers and Special Combinatorial Sequences

In this section, we use moment forms of special combinatorial sequences, characteristic function and generating function method to investigate the relationships between Daehee numbers D_n , Cauchy numbers of the second kind \hat{C}_n , derangement numbers d_n , and Stirling numbers of the first kind, then we obtain combinatorial identities about them.

Theorem 9. Let $r.v \ u \sim U[0,1], \ \Gamma \sim \Gamma(1,1), \ X \sim \Gamma(u,1),$ then Daehee numbers D_n , Cauchy numbers of the second kind \hat{C}_n , and derangement numbers d_n satisfy the following identity

$$\sum_{k=0}^{n} \binom{n}{k} (-1)^n D_k \hat{C}_{n-k} = \sum_{k=0}^{n} \binom{n}{k} d_k, n \geqslant 0.$$
 (36)

Proof. On one hand,

$$[E(u\Gamma) - EX]^n = \sum_{k=0}^n \binom{n}{k} E(u\Gamma)^k (-EX)^{n-k}$$
$$= \sum_{k=0}^n \binom{n}{k} (-1)^n E(-u\Gamma)^k EX^{n-k}$$

$$= \sum_{k=0}^{n} \binom{n}{k} (-1)^n D_k \hat{C}_{n-k}, \tag{37}$$

Write the generating function of the equation above, noting that $i^2 = -1$

$$\sum_{n=0}^{\infty} [E(u\Gamma) - EX]^n \frac{(it)^n}{n!} = \sum_{n=0}^{\infty} \sum_{k=0}^n \binom{n}{k} (-1)^n D_k \hat{C}_{n-k} \frac{(it)^n}{n!}$$

$$= \sum_{n=0}^{\infty} D_n \frac{(-it)^n}{n!} \sum_{n=0}^{\infty} \hat{C}_n \frac{(-it)^n}{n!}$$

$$= \frac{\ln(1-it)}{-it} \frac{-it}{(1-it)\ln(1-it)}$$

$$= Ee^{it\Gamma} = \sum_{n=0}^{\infty} E\Gamma^n \frac{(it)^n}{n!},$$
(38)

On the other hand, by comparing the coefficients of $\frac{(it)^n}{n!}$ in Eq.(38), we have

$$[E(u\Gamma) - EX]^n = E\Gamma^n = E(\Gamma - 1 + 1)^n = \sum_{k=0}^n \binom{n}{k} E(\Gamma - 1)^k = \sum_{k=0}^n \binom{n}{k} d_k.$$
 (39)

From Eq.(37) and Eq.(39), we can get Eq.(36). Theorem 9 is proved.

Theorem 10. Higher-order Deahee numbers $D_n^{(k)}$ and Stirling numbers of the first kind satisfy the following relationship

$$D_n^{(k)} = \frac{s(n+k,k)}{\binom{n+k}{k}}. (40)$$

Proof. From corollary 1 and lemma 3, write generating function of the left-hand side of Eq.(40),

$$\sum_{n=0}^{\infty} D_n^{(k)} \frac{t^n}{n!} = \sum_{n=0}^{\infty} (-1)^n E(u_1 \Gamma_1 + \dots + u_k \Gamma_k)^n \frac{t^n}{n!}$$

$$= \sum_{n=0}^{\infty} (-1)^n \binom{n+k}{k} E(u_1 \Gamma_1 + \dots + u_k \Gamma_k)^n \frac{k!}{t^k} \frac{t^{n+k}}{(n+k)!}$$

$$= \sum_{n=0}^{\infty} \frac{s(n+k,k)}{\binom{n+k}{k}} \frac{t^n}{n!}.$$

By comparing the coefficients of $\frac{t^n}{n!}$, we obtain the conclusion.

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Acknowledgements

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